

The Sample Complexity (with a Generative Model)

Sham Kakade and Kianté Brantley

CS 2824: Foundations of Reinforcement Learning

Announcements

- **HW1** is posted now!
 - it cover many concepts from class
- **Reading assignments**
 - please do the readings; **they are helpful for our lectures**
 - next reading assignment posted soon

Today:

Today:

- Recap: computational complexity
 - Question: Given an MDP $\mathcal{M} = (S, A, P, r, \gamma)$ can we **exactly compute** Q^* (or find π^*) in polynomial time?

Today:

- Recap: computational complexity
 - Question: Given an MDP $\mathcal{M} = (S, A, P, r, \gamma)$ can we **exactly compute** Q^* (or find π^*) in polynomial time?
- Today: **statistical complexity**
 - Question: Given only sampling access to an unknown MDP $\mathcal{M} = (S, A, P, r, \gamma)$ how many **observed transitions do we need** to **estimate** Q^* (or find π^*)?

Recap

Summary Table

	Value Iteration	Policy Iteration	LP-based Algorithms
Poly.	$S^2 A \frac{L(P,r,\gamma) \log \frac{1}{1-\gamma}}{1-\gamma}$	$(S^3 + S^2 A) \frac{L(P,r,\gamma) \log \frac{1}{1-\gamma}}{1-\gamma}$	$S^3 A L(P, r, \gamma)$
Strongly Poly.	X	$(S^3 + S^2 A) \cdot \min \left\{ \frac{A^S}{S}, \frac{S^2 A \log \frac{S^2}{1-\gamma}}{1-\gamma} \right\}$	$S^4 A^4 \log \frac{S}{1-\gamma}$

- VI: poly time for **fixed** γ , not strongly poly
- PI: poly and strongly-poly time for **fixed** γ
- LP approach: poly and strongly-poly time
(LP approach is only logarithmic in $1/(1 - \gamma)$)

Today

Sampling Models

(for learning in an unknown MDP)

- **Episodic setting:**
 - in every episode, $s_0 \sim \mu$.
 - the learner acts for some finite number of steps and observes the trajectory.
 - The state then resets to $s_0 \sim \mu$.

Sampling Models

(for learning in an unknown MDP)

- **Episodic setting:**
 - in every episode, $s_0 \sim \mu$.
 - the learner acts for some finite number of steps and observes the trajectory.
 - The state then resets to $s_0 \sim \mu$.
- **Offline RL setting:**
 - You have a collection of observed transitions/reward: $\{(s, a, s', r(s, a))\}$
 - You don't have control of the data generating distribution.

Sampling Models

(for learning in an unknown MDP)

- **Episodic setting:**
 - in every episode, $s_0 \sim \mu$.
 - the learner acts for some finite number of steps and observes the trajectory.
 - The state then resets to $s_0 \sim \mu$.
- **Offline RL setting:**
 - You have a collection of observed transitions/reward: $\{(s, a, s', r(s, a))\}$
 - You don't have control of the data generating distribution.
- **Generative model setting:**
 - input: (s, a) output: a sample $s' \sim P(\cdot | s, a)$ and $r(s, a)$
 - provides an “idealized model” to study statistical limits

Sampling Models

(for learning in an unknown MDP)

- **Episodic setting:**
 - in every episode, $s_0 \sim \mu$.
 - the learner acts for some finite number of steps and observes the trajectory.
 - The state then resets to $s_0 \sim \mu$.
- **Offline RL setting:**
 - You have a collection of observed transitions/reward: $\{(s, a, s', r(s, a))\}$
 - You don't have control of the data generating distribution.
- **Generative model setting:**
 - input: (s, a) output: a sample $s' \sim P(\cdot | s, a)$ and $r(s, a)$
 - provides an “idealized model” to study statistical limits
- **Sample complexity of RL:**
how many transitions do we need observe in order to find a near optimal policy?

Two Fundamental Questions in Sample Complexity

1. The Model Size Question (Sublinear Learning)

- How many parameters do we need to specify the transition kernel P ?
(and how many for the policy?)
- Q1: Can we find an ϵ -optimal policy with **sublinear** sample complexity?

Two Fundamental Questions in Sample Complexity

1. The Model Size Question (Sublinear Learning)

- How many parameters do we need to specify the transition kernel P ? (and how many for the policy?)
- Q1: Can we find an ϵ -optimal policy with **sublinear** sample complexity?

2. The Horizon Question (Error Amplification)

- The Scale: In discounted settings, the values scale as $1/(1 - \gamma)$.
- Target: It is natural to measure our additive error ϵ relative to this scale.
- Q2: What is the “**horizon amplification**”? i.e. the dependence on $1/(1 - \gamma)$

Attempt 1:
the naive model based approach

The most naive approach: model based

- Today: let us assume access to a [generative model](#)

The most naive approach: model based

- Today: let us assume access to a **generative model**
- most naive approach to learning:
 - Call our simulator **N times at each state action pair.**
 - Let \hat{P} be our empirical model:

$$\hat{P}(s' | s, a) = \frac{\text{count}(s', s, a)}{N}$$

where $\text{count}(s', s, a)$ is the #times (s, a) transitions to state s' .

- we know the reward after one call.
(for simplicity, we often assume $r(s, a)$ is deterministic)

The most naive approach: model based

- Today: let us assume access to a **generative model**
- most naive approach to learning:
 - Call our simulator **N times at each state action pair.**
 - Let \hat{P} be our empirical model:

$$\hat{P}(s' | s, a) = \frac{\text{count}(s', s, a)}{N}$$

where $\text{count}(s', s, a)$ is the #times (s, a) transitions to state s' .

- we know the reward after one call.
(for simplicity, we often assume $r(s, a)$ is deterministic)
- The total number of calls to our generative model is **SAN**.

Model accuracy

Proposition: c is an absolute constant. $\epsilon > 0$. For $N \geq \frac{c\gamma}{(1-\gamma)^4} \frac{S \log(cSA/\delta)}{\epsilon^2}$
and with probability greater than $1 - \delta$,

Model accuracy

Proposition: c is an absolute constant. $\epsilon > 0$. For $N \geq \frac{c\gamma}{(1-\gamma)^4} \frac{S \log(cSA/\delta)}{\epsilon^2}$

and with probability greater than $1 - \delta$,

- Model accuracy: The transition model is ϵ has error bounded as:

$$\max_{s,a} \|P(\cdot | s, a) - \widehat{P}(\cdot | s, a)\|_1 \leq (1 - \gamma)^2 \epsilon / 2.$$

Model accuracy

Proposition: c is an absolute constant. $\epsilon > 0$. For $N \geq \frac{c\gamma}{(1-\gamma)^4} \frac{S \log(cSA/\delta)}{\epsilon^2}$

and with probability greater than $1 - \delta$,

- Model accuracy: The transition model is ϵ has error bounded as:

$$\max_{s,a} \|P(\cdot | s, a) - \widehat{P}(\cdot | s, a)\|_1 \leq (1 - \gamma)^2 \epsilon / 2.$$

- Uniform value accuracy: For all policies π ,

$$\|Q^\pi - \widehat{Q}^\pi\|_\infty \leq \epsilon / 2$$

Model accuracy

Proposition: c is an absolute constant. $\epsilon > 0$. For $N \geq \frac{c\gamma}{(1-\gamma)^4} \frac{S \log(cSA/\delta)}{\epsilon^2}$

and with probability greater than $1 - \delta$,

- Model accuracy: The transition model is ϵ has error bounded as:

$$\max_{s,a} \|P(\cdot | s, a) - \widehat{P}(\cdot | s, a)\|_1 \leq (1 - \gamma)^2 \epsilon / 2.$$

- Uniform value accuracy: For all policies π ,

$$\|Q^\pi - \widehat{Q}^\pi\|_\infty \leq \epsilon / 2$$

- Near optimal planning: Suppose that $\widehat{\pi}^*$ is the optimal policy in \widehat{M} .

$$\|Q^* - Q^{\widehat{\pi}^*}\|_\infty \leq \epsilon$$

Matrix Expressions

- View P as a matrix of size $SA \times S$ (rows are probability distributions)

Matrix Expressions

- View P as a matrix of size $SA \times S$ (rows are probability distributions)
- Define P^π to be the transition matrix on state-action pairs (for deterministic π):

$$P_{(s,a),(s',a')}^\pi := \begin{cases} P(s' | s, a) & \text{if } a' = \pi(s') \\ 0 & \text{if } a' \neq \pi(s') \end{cases}$$

Matrix Expressions

- View P as a matrix of size $SA \times S$ (rows are probability distributions)
- Define P^π to be the transition matrix on state-action pairs (for deterministic π):

$$P^\pi_{(s,a),(s',a')} := \begin{cases} P(s' | s, a) & \text{if } a' = \pi(s') \\ 0 & \text{if } a' \neq \pi(s') \end{cases}$$

- With this notation,
 $Q^\pi = r + \gamma P V^\pi$
 $Q^\pi = r + \gamma P^\pi Q^\pi$

Matrix Expressions

- View P as a matrix of size $SA \times S$ (rows are probability distributions)
- Define P^π to be the transition matrix on state-action pairs (for deterministic π):

$$P^\pi_{(s,a),(s',a')} := \begin{cases} P(s' | s, a) & \text{if } a' = \pi(s') \\ 0 & \text{if } a' \neq \pi(s') \end{cases}$$

- With this notation,

$$Q^\pi = r + \gamma P V^\pi$$

$$Q^\pi = r + \gamma P^\pi Q^\pi$$

- Also,

$$Q^\pi = (I - \gamma P^\pi)^{-1} r$$

(where one can show the inverse exists)

“Simulation” Lemma

“Simulation Lemma”: For all π ,

$$Q^\pi - \widehat{Q}^\pi = \gamma(I - \gamma \widehat{P}^\pi)^{-1}(P - \widehat{P})V^\pi$$

“Simulation” Lemma

“Simulation Lemma”: For all π ,

$$Q^\pi - \widehat{Q}^\pi = \gamma(I - \gamma \widehat{P}^\pi)^{-1}(P - \widehat{P})V^\pi$$

Proof: Using our matrix equality for Q^π , we have:

$$\begin{aligned} Q^\pi - \widehat{Q}^\pi &= Q^\pi - (I - \gamma \widehat{P}^\pi)^{-1}r \\ &= (I - \gamma \widehat{P}^\pi)^{-1}((I - \gamma \widehat{P}^\pi) - (I - \gamma P^\pi))Q^\pi \\ &= \gamma(I - \gamma \widehat{P}^\pi)^{-1}(P^\pi - \widehat{P}^\pi)Q^\pi \\ &= \gamma(I - \gamma \widehat{P}^\pi)^{-1}(P - \widehat{P})V^\pi \end{aligned}$$

Proof of Claim 1

Proof of Claim 1

- Concentration of a distribution in the ℓ_1 norm:
 - For a fixed s, a . With pr greater than $1 - \delta$,

$$\|P(\cdot | s, a) - \hat{P}(\cdot | s, a)\|_1 \leq c \sqrt{\frac{S \log(1/\delta)}{N}}$$

with N samples used to estimate $\hat{P}(\cdot | s, a)$.

Proof of Claim 1

- Concentration of a distribution in the ℓ_1 norm:

- For a fixed s, a . With pr greater than $1 - \delta$,

$$\|P(\cdot | s, a) - \widehat{P}(\cdot | s, a)\|_1 \leq c \sqrt{\frac{S \log(1/\delta)}{N}}$$

with N samples used to estimate $\widehat{P}(\cdot | s, a)$.

- The first claim now follows by the union bound.

Proof of Claim 2 (&3)

Proof of Claim 2 (&3)

For the second claim,

$$\|Q^\pi - \widehat{Q}^\pi\|_\infty = \|\gamma(I - \gamma \widehat{P}^\pi)^{-1}(P - \widehat{P})V^\pi\|_\infty$$

Proof of Claim 2 (&3)

For the second claim,

$$\|Q^\pi - \widehat{Q}^\pi\|_\infty = \|\gamma(I - \gamma \widehat{P}^\pi)^{-1}(P - \widehat{P})V^\pi\|_\infty$$

$$\leq \frac{\gamma}{1 - \gamma} \|(P - \widehat{P})V^\pi\|_\infty$$

$$\leq \frac{\gamma}{1 - \gamma} \left(\max_{s,a} \|P(\cdot | s, a) - \widehat{P}(\cdot | s, a)\|_1 \right) \|V^\pi\|_\infty$$

$$\leq \frac{\gamma}{(1 - \gamma)^2} \max_{s,a} \|P(\cdot | s, a) - \widehat{P}(\cdot | s, a)\|_1$$

(why is the first inequality true?)

Proof of Claim 2 (&3)

For the second claim,

$$\|Q^\pi - \widehat{Q}^\pi\|_\infty = \|\gamma(I - \gamma \widehat{P}^\pi)^{-1}(P - \widehat{P})V^\pi\|_\infty$$

$$\leq \frac{\gamma}{1 - \gamma} \|(P - \widehat{P})V^\pi\|_\infty$$

$$\leq \frac{\gamma}{1 - \gamma} \left(\max_{s,a} \|P(\cdot | s, a) - \widehat{P}(\cdot | s, a)\|_1 \right) \|V^\pi\|_\infty$$

$$\leq \frac{\gamma}{(1 - \gamma)^2} \max_{s,a} \|P(\cdot | s, a) - \widehat{P}(\cdot | s, a)\|_1$$

(why is the first inequality true?)

The proof for the Claim 3 immediately follows from the second claim.

Attempt 2:

obtaining sublinear sample complexity

idea: use concentration only on V^*

Reference sheet (defs/notation)

Reference sheet (defs/notation)

- Remember: # samples from generative model = SAN

Reference sheet (defs/notation)

- Remember: # samples from generative model = SAN
- P^π is the transition matrix on state-action pairs for a deterministic policy π :

$$P_{(s,a),(s',a')}^\pi := P(s' | s, a) \quad \text{if } a' = \pi(s')$$
$$0 \quad \text{if } a' \neq \pi(s')$$

Reference sheet (defs/notation)

- Remember: # samples from generative model = SAN
- P^π is the transition matrix on state-action pairs for a deterministic policy π :

$$P_{(s,a),(s',a')}^\pi := P(s' | s, a) \quad \text{if } a' = \pi(s')$$
$$0 \quad \text{if } a' \neq \pi(s')$$

- With this notation,

$$Q^\pi = r + \gamma P V^\pi, \quad Q^\pi = r + \gamma P^\pi Q^\pi, \quad Q^\pi = (I - \gamma P^\pi)^{-1} r$$

Reference sheet (defs/notation)

- Remember: # samples from generative model = SAN

- P^π is the transition matrix on state-action pairs for a deterministic policy π :

$$P_{(s,a),(s',a')}^\pi := P(s' | s, a) \quad \text{if } a' = \pi(s')$$
$$0 \quad \text{if } a' \neq \pi(s')$$

- With this notation,

$$Q^\pi = r + \gamma P V^\pi, \quad Q^\pi = r + \gamma P^\pi Q^\pi, \quad Q^\pi = (I - \gamma P^\pi)^{-1} r$$

- $\frac{1}{1-\gamma}(I - \gamma P^\pi)^{-1}$ is a matrix whose rows are probability distributions (why?)

Reference sheet (defs/notation)

- Remember: # samples from generative model = SAN
- P^π is the transition matrix on state-action pairs for a deterministic policy π :
$$P_{(s,a),(s',a')}^\pi := P(s' | s, a) \quad \text{if } a' = \pi(s')$$
$$0 \quad \text{if } a' \neq \pi(s')$$
- With this notation,
$$Q^\pi = r + \gamma P V^\pi, \quad Q^\pi = r + \gamma P^\pi Q^\pi, \quad Q^\pi = (I - \gamma P^\pi)^{-1} r$$
- $\frac{1}{1 - \gamma} (I - \gamma P^\pi)^{-1}$ is a matrix whose rows are probability distributions (why?)
- \widehat{Q}^* : optimal value in estimated model \widehat{M} .
 $\widehat{\pi}^*$: optimal policy in \widehat{M} .
 $Q^{\widehat{\pi}^*}$: (true) value of estimated policy.

Attempt 2: Sublinear Sample Complexity

Attempt 2: Sublinear Sample Complexity

Proposition: (Crude Value Bound) With probability greater than $1 - \delta$,

$$\|Q^* - \widehat{Q}^*\|_\infty \leq \frac{\gamma}{(1 - \gamma)^2} \sqrt{\frac{2 \log(2SA/\delta)}{N}}$$

$$\|Q^* - \widehat{Q}^{\pi^*}\|_\infty \leq \frac{\gamma}{(1 - \gamma)^2} \sqrt{\frac{2 \log(2SA/\delta)}{N}}$$

Attempt 2: Sublinear Sample Complexity

Proposition: (Crude Value Bound) With probability greater than $1 - \delta$,

$$\|Q^* - \widehat{Q}^*\|_\infty \leq \frac{\gamma}{(1-\gamma)^2} \sqrt{\frac{2 \log(2SA/\delta)}{N}}$$

$$\|Q^* - \widehat{Q}^{\pi^*}\|_\infty \leq \frac{\gamma}{(1-\gamma)^2} \sqrt{\frac{2 \log(2SA/\delta)}{N}}$$

What about the value of the policy?

$$\|Q^* - Q^{\widehat{\pi}^*}\|_\infty \leq \frac{\gamma}{(1-\gamma)^3} \sqrt{\frac{2 \log(2SA/\delta)}{N}}$$

Sample Size Corollaries

Corollary: for $\epsilon < 1$, provided $N \geq \frac{c}{(1-\gamma)^4} \frac{\log(cSA/\delta)}{\epsilon^2}$ then

$$\|Q^* - \widehat{Q}^*\|_\infty \leq \epsilon \text{ (with prob. greater than } 1 - \delta)$$

Sample Size Corollaries

Corollary: for $\epsilon < 1$, provided $N \geq \frac{c}{(1-\gamma)^4} \frac{\log(cSA/\delta)}{\epsilon^2}$ then

$$\|Q^* - \widehat{Q}^*\|_\infty \leq \epsilon \text{ (with prob. greater than } 1 - \delta)$$

What about the policy?

Corollary: for $\epsilon < 1$, provided $N \geq \frac{c}{(1-\gamma)^6} \frac{\log(cSA/\delta)}{\epsilon^2}$ then

$$\|Q^* - Q^{\widehat{\pi}^*}\|_\infty \leq \epsilon \text{ (with prob. greater than } 1 - \delta)$$

Component-wise Bounds Lemma

Lemma: we have that

$$Q^* - \widehat{Q}^* \leq \gamma(I - \gamma \widehat{P}^{\pi^*})^{-1}(P - \widehat{P})V^*$$

$$Q^* - \widehat{Q}^* \geq \gamma(I - \gamma \widehat{P}^{\hat{\pi}^*})^{-1}(P - \widehat{P})V^*$$

Component-wise Bounds Lemma

Lemma: we have that

$$Q^* - \widehat{Q}^* \leq \gamma(I - \gamma \widehat{P}^{\pi^*})^{-1}(P - \widehat{P})V^*$$

$$Q^* - \widehat{Q}^* \geq \gamma(I - \gamma \widehat{P}^{\hat{\pi}^*})^{-1}(P - \widehat{P})V^*$$

Proof:

For the first claim, the optimality of π^* in M implies:

$$Q^* - \widehat{Q}^* = Q^{\pi^*} - \widehat{Q}^{\hat{\pi}^*} \leq Q^{\pi^*} - \widehat{Q}^{\pi^*} = \gamma(I - \gamma \widehat{P}^{\pi^*})^{-1}(P - \widehat{P})V^*,$$

using the simulation lemma in the final step.

See notes for the proof of second claim.

Proof: (& key idea for sublinearity)

Proof: (& key idea for sublinearity)

- Proof of the first claim:

Proof: (& key idea for sublinearity)

- Proof of the first claim:

- By comp. lemma: $\|Q^\star - \widehat{Q}^\star\|_\infty \leq \frac{\gamma}{1-\gamma} \|(P - \widehat{P})V^\star\|_\infty$

Proof: (& key idea for sublinearity)

- Proof of the first claim:

- By comp. lemma: $\|Q^\star - \widehat{Q}^\star\|_\infty \leq \frac{\gamma}{1-\gamma} \|(P - \widehat{P})V^\star\|_\infty$

- Recall $\|V^\star\|_\infty \leq 1/(1-\gamma)$.

Proof: (& key idea for sublinearity)

- Proof of the first claim:

- By comp. lemma: $\|Q^\star - \widehat{Q}^\star\|_\infty \leq \frac{\gamma}{1-\gamma} \|(P - \widehat{P})V^\star\|_\infty$

- Recall $\|V^\star\|_\infty \leq 1/(1-\gamma)$.

- By Hoeffding's inequality and the union bound,

$$\begin{aligned} \|(P - \widehat{P})V^\star\|_\infty &= \max_{s,a} \left| E_{s' \sim P(\cdot|s,a)}[V^\star(s')] - E_{s' \sim \widehat{P}(\cdot|s,a)}[V^\star(s')] \right| \\ &\leq \frac{1}{1-\gamma} \sqrt{\frac{2 \log(2SA/\delta)}{N}} \end{aligned}$$

which holds with probability greater than $1 - \delta$.

Proof: (& key idea for sublinearity)

- Proof of the first claim:

- By comp. lemma: $\|Q^* - \widehat{Q}^*\|_\infty \leq \frac{\gamma}{1-\gamma} \|(P - \widehat{P})V^*\|_\infty$

- Recall $\|V^*\|_\infty \leq 1/(1-\gamma)$.

- By Hoeffding's inequality and the union bound,

$$\begin{aligned} \|(P - \widehat{P})V^*\|_\infty &= \max_{s,a} \left| E_{s' \sim P(\cdot|s,a)}[V^*(s')] - E_{s' \sim \widehat{P}(\cdot|s,a)}[V^*(s')] \right| \\ &\leq \frac{1}{1-\gamma} \sqrt{\frac{2 \log(2SA/\delta)}{N}} \end{aligned}$$

which holds with probability greater than $1 - \delta$.

- Proof of second claim is similar (see the book)

Attempt 3:

minimax optimal sample complexity

idea: better variance control

(“near”) Minimax Optimal Sample Complexity

(“near”) Minimax Optimal Sample Complexity

Theorem: With probability greater than $1 - \delta$,

$$\|Q^* - \widehat{Q}^*\|_\infty \leq \gamma \sqrt{\frac{c}{(1-\gamma)^3} \frac{\log(cSA/\delta)}{N}} + \frac{c\gamma}{(1-\gamma)^3} \frac{\log(cSA/\delta)}{N},$$

where c is an absolute constant.

Minimax Optimal Sample Complexity

Minimax Optimal Sample Complexity

Corollary: for $\epsilon < 1$, provided $N \geq \frac{c}{(1-\gamma)^3} \frac{\log(cSA/\delta)}{\epsilon^2}$ then

$$\|Q^* - \widehat{Q}^*\|_\infty \leq \epsilon \text{ (with prob. greater than } 1 - \delta)$$

Minimax Optimal Sample Complexity

Corollary: for $\epsilon < 1$, provided $N \geq \frac{c}{(1-\gamma)^3} \frac{\log(cSA/\delta)}{\epsilon^2}$ then

$$\|Q^* - \widehat{Q}^*\|_\infty \leq \epsilon \text{ (with prob. greater than } 1 - \delta)$$

What about the policy?

Naively, we need $N/(1-\gamma)^2$ more samples.

A different thm gives: With the same N ,

$$\|Q^* - Q^{\widehat{\pi}^*}\|_\infty \leq \epsilon \text{ (with prob. greater than } 1 - \delta)$$

Minimax Optimal Sample Complexity

Corollary: for $\epsilon < 1$, provided $N \geq \frac{c}{(1-\gamma)^3} \frac{\log(cSA/\delta)}{\epsilon^2}$ then

$$\|Q^* - \widehat{Q}^*\|_\infty \leq \epsilon \text{ (with prob. greater than } 1 - \delta)$$

What about the policy?

Naively, we need $N/(1-\gamma)^2$ more samples.

A different thm gives: With the same N ,

$$\|Q^* - Q^{\widehat{\pi}^*}\|_\infty \leq \epsilon \text{ (with prob. greater than } 1 - \delta)$$

Lower Bound: We can't do better.

Revisiting proof attempt 2: where is there slop?

- Proof of the first claim:

- By comp. lemma: $\|Q^* - \widehat{Q}^*\|_\infty \leq \frac{\gamma}{1 - \gamma} \|(P - \widehat{P})V^*\|_\infty$

- Recall $\|V^*\|_\infty \leq 1/(1 - \gamma)$.

- By Hoeffding's inequality and the union bound,

$$\begin{aligned} \|(P - \widehat{P})V^*\|_\infty &= \max_{s,a} \left| E_{s' \sim P(\cdot|s,a)}[V^*(s')] - E_{s' \sim \widehat{P}(\cdot|s,a)}[V^*(s')] \right| \\ &\leq \frac{1}{1 - \gamma} \sqrt{\frac{2 \log(2SA/\delta)}{N}} \end{aligned}$$

which holds with probability greater than $1 - \delta$.

- Proof of second claim is similar (see the book)

Proof sketch: part 1

- From Bernstein's ineq, with pr. greater than $1 - \delta$, we have (component-wise):

$$|(P - \hat{P})V^*| \leq \sqrt{\frac{2 \log(2SA/\delta)}{N}} \sqrt{\text{Var}_P(V^*)} + \frac{1}{1 - \gamma} \frac{2 \log(2SA/\delta)}{3N} \vec{1}$$

Proof sketch: part 1

- From Bernstein's ineq, with pr. greater than $1 - \delta$, we have (component-wise):

$$|(P - \widehat{P})V^*| \leq \sqrt{\frac{2 \log(2SA/\delta)}{N}} \sqrt{\text{Var}_P(V^*)} + \frac{1}{1 - \gamma} \frac{2 \log(2SA/\delta)}{3N} \vec{1}$$

- How to use this: again from “Component-wise Bounds” lemma,

$$Q^* - \widehat{Q}^* \leq \gamma \|(I - \gamma \widehat{P}^{\pi^*})^{-1} (P - \widehat{P})V^*\|_\infty \leq ??$$

Proof sketch: part 1

- From Bernstein's ineq, with pr. greater than $1 - \delta$, we have (component-wise):

$$|(P - \widehat{P})V^*| \leq \sqrt{\frac{2 \log(2SA/\delta)}{N}} \sqrt{\text{Var}_P(V^*)} + \frac{1}{1 - \gamma} \frac{2 \log(2SA/\delta)}{3N} \vec{1}$$

- How to use this: again from “Component-wise Bounds” lemma,

$$Q^* - \widehat{Q}^* \leq \gamma \|(I - \gamma \widehat{P}^{\pi^*})^{-1} (P - \widehat{P})V^*\|_\infty \leq ??$$

- Therefore

$$Q^* - \widehat{Q}^* \leq \gamma \sqrt{\frac{2 \log(2SA/\delta)}{N}} \|(I - \gamma \widehat{P}^{\pi^*})^{-1} \sqrt{\text{Var}_P(V^*)}\|_\infty$$

+ "lower order term"

Bellman Equation for the (total) Variance

- **Variance:** $\text{Var}_P(V)(s, a) := \text{Var}_{P(\cdot|s,a)}(V)$

Component wise variance: $\text{Var}_P(V) := P(V)^2 - (PV)^2$

Bellman Equation for the (total) Variance

- **Variance:** $\text{Var}_P(V)(s, a) := \text{Var}_{P(\cdot|s,a)}(V)$

Component wise variance: $\text{Var}_P(V) := P(V)^2 - (PV)^2$

- Let's keep around the MDP M subscripts.

Define Σ_M^π as the (total) variance of the discounted reward:

$$\Sigma_M^\pi(s, a) := E \left[\left(\sum_{t=0}^{\infty} \gamma^t r(s_t, a_t) - Q_M^\pi(s, a) \right)^2 \middle| s_0 = s, a_0 = a \right]$$

Bellman Equation for the (total) Variance

- **Variance:** $\text{Var}_P(V)(s, a) := \text{Var}_{P(\cdot|s,a)}(V)$

Component wise variance: $\text{Var}_P(V) := P(V)^2 - (PV)^2$

- Let's keep around the MDP M subscripts.

Define Σ_M^π as the (total) variance of the discounted reward:

$$\Sigma_M^\pi(s, a) := E \left[\left(\sum_{t=0}^{\infty} \gamma^t r(s_t, a_t) - Q_M^\pi(s, a) \right)^2 \middle| s_0 = s, a_0 = a \right]$$

- **Bellman equation for the total variance:**

$$\Sigma_M^\pi = \gamma^2 \text{Var}_P(V_M^\pi) + \gamma^2 P^\pi \Sigma_M^\pi$$

Key Lemma

Lemma: For any policy π and MDP M ,

$$\left\| (I - \gamma P^\pi)^{-1} \sqrt{\text{Var}_P(V_M^\pi)} \right\|_\infty \leq \sqrt{\frac{2}{(1 - \gamma)^3}}$$

Proof idea: convexity + Bellman equations for the variance.

Putting it all together

Proof sketch: we have two MDPs M and \hat{M} . need to bound:

Putting it all together

Proof sketch: we have two MDPs M and \widehat{M} . need to bound:

$$\|(I - \gamma \widehat{P}^{\pi^*})^{-1} \sqrt{\text{Var}_P(V^*)}\|_\infty = \|(I - \gamma P_{\widehat{M}}^{\pi^*})^{-1} \sqrt{\text{Var}_P(V_M^{\pi^*})}\|_\infty$$

$$\leq \|(I - \gamma P_{\widehat{M}}^{\pi^*})^{-1} \sqrt{\text{Var}_P(V_{\widehat{M}}^{\pi^*})}\| + \text{"lower order"}$$

$$\leq \sqrt{\frac{2}{(1 - \gamma)^3}} + \text{"lower order"}$$

Putting it all together

Proof sketch: we have two MDPs M and \widehat{M} . need to bound:

$$\|(I - \gamma \widehat{P}^{\pi^*})^{-1} \sqrt{\text{Var}_P(V^*)}\|_\infty = \|(I - \gamma P_{\widehat{M}}^{\pi^*})^{-1} \sqrt{\text{Var}_P(V_M^{\pi^*})}\|_\infty$$

$$\leq \|(I - \gamma P_{\widehat{M}}^{\pi^*})^{-1} \sqrt{\text{Var}_P(V_{\widehat{M}}^{\pi^*})}\|_\infty + \text{"lower order"}$$

$$\leq \sqrt{\frac{2}{(1 - \gamma)^3}} + \text{"lower order"}$$

First equality above: just notation

Putting it all together

Proof sketch: we have two MDPs M and \widehat{M} . need to bound:

$$\|(I - \gamma \widehat{P}^{\pi^*})^{-1} \sqrt{\text{Var}_P(V^*)}\|_\infty = \|(I - \gamma P_{\widehat{M}}^{\pi^*})^{-1} \sqrt{\text{Var}_P(V_M^{\pi^*})}\|_\infty$$

$$\leq \|(I - \gamma P_{\widehat{M}}^{\pi^*})^{-1} \sqrt{\text{Var}_P(V_{\widehat{M}}^{\pi^*})}\|_\infty + \text{"lower order"}$$

$$\leq \sqrt{\frac{2}{(1 - \gamma)^3}} + \text{"lower order"}$$

First equality above: just notation

Second step: concentration \rightarrow we need to quantify:

$$\sqrt{\text{Var}_P(V_M^{\pi^*})} \approx \sqrt{\text{Var}_P(V_{\widehat{M}}^{\pi^*})}$$

Putting it all together

Proof sketch: we have two MDPs M and \widehat{M} . need to bound:

$$\|(I - \gamma \widehat{P}^{\pi^*})^{-1} \sqrt{\text{Var}_P(V^*)}\|_\infty = \|(I - \gamma P_{\widehat{M}}^{\pi^*})^{-1} \sqrt{\text{Var}_P(V_M^{\pi^*})}\|_\infty$$

$$\leq \|(I - \gamma P_{\widehat{M}}^{\pi^*})^{-1} \sqrt{\text{Var}_P(V_{\widehat{M}}^{\pi^*})}\|_\infty + \text{"lower order"}$$

$$\leq \sqrt{\frac{2}{(1 - \gamma)^3}} + \text{"lower order"}$$

First equality above: just notation

Second step: concentration \rightarrow we need to quantify:

$$\sqrt{\text{Var}_P(V_M^{\pi^*})} \approx \sqrt{\text{Var}_P(V_{\widehat{M}}^{\pi^*})}$$

Last step: previous slide